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| 4              | Lymph node counts and hospital survival rates after resection for colon cancer  |
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26

## 27 Abstract

| 28 | Context. Several studies suggest improved survival among patients in whom a higher                    |
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| 29 | number of nodes are examined after colectomy for colon cancer. The National Quality                   |
| 30 | Forum and other organizations recently endorsed a 12-node minimum as a measure of                     |
| 31 | hospital quality.   |
| 32 | Objective. To assess whether hospitals that examine more lymph nodes after resection for              |
| 33 | colon cancer have superior late survival rates.   |
| 34 | Design. Using the national Surveillance Epidemiology and End Results (SEER)-Medicare                  |
| 35 | linked database (1995-2005), we first identified all patients undergoing colectomy for non-           |
| 36 | metastatic colon cancer ( <b>n=30,625</b> ). We first ranked hospitals according to the proportion of |
| 37 | their patients in whom 12 or more lymph nodes were examined and then sorted them into 4               |
| 38 | evenly sized groups. We then assessed late survival rates for each hospital group, adjusting for      |
| 39 | potentially confounding patient and provider characteristics.   |
| 40 | Study Participants. US Medicare patients residing in SEER regions.                                    |
| 41 | Main Outcome Measures. 5-year survival  |
| 42 | <u>Results.</u> Hospitals with the highest proportions of patients with 12 or more lymph nodes        |
| 43 | tended to treat lower risk patients and had substantially high procedure volumes. After               |
| 44 | adjusting for these and other factors, there remained no statistically significant relationship       |
| 45 | between hospital lymph node examination rates and survival after surgery (adjusted hazard             |

- 46 ratio, high vs. low hospital quartile, 0.95, 95% CI 0.88-1.03). Hospital lymph node
- 47 examination rates were not related to the overall use of adjuvant chemotherapy (26% vs. 25%,
- 48 high vs. low hospital quartile).

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- 49 <u>Conclusions.</u> Lymph node examination rates do not predict hospital survival rates
- 50 following colectomy for colon cancer. Efforts by payers and professional organizations to
- 51 increase node counts with this quality indicator may have limited value as a public health
- 52 intervention.

#### 53 Introduction

54 It may be important that a sufficient number of lymph nodes are obtained and 55 examined at the time of primary resection for colon cancer. More complete node 56 clearance may itself result in lower rates of local or distant cancer recurrence. Obtaining 57 more lymph nodes may also benefit patients to the extent that it allows for more accurate 58 cancer staging and thus more appropriate use of adjuvant chemotherapy for patients with 59 node-positive disease (stage III cancer). Numerous observational studies and a recent 60 systematic review suggest that patients in whom a high number of nodes are examined 61 have considerably lower late mortality after colectomy for colon cancer than patients with fewer nodes examined.<sup>1-5</sup> Such studies have prompted interest in using minimum lymph 62 63 node counts as a quality indicator for colon cancer resection. Recently, in collaboration 64 with the American College of Surgeons, the American Society for Clinical Oncology, 65 the National Comprehensive Cancer Network, and other stakeholders, the National 66 Quality Forum (NQF) endorsed a 12-node minimum as a consensus standard for hospital-based performance with colectomy for colon cancer.<sup>6</sup> Large private payers 67 68 have already begun incorporating this measure into its pay-for-performance programs.<sup>7</sup> 69

70 Whether such efforts will ultimately improve patient outcomes with colon 71 cancer remain unclear. Apparent associations between the number of lymph nodes 72 examined and survival after colon resection may reflect confounding patient factors as 73 much as quality of care. There is wide biological variation in the quantity and distribution 74 of mesenteric lymph nodes among patients. <sup>8</sup> Patients with more nodes may have a better 75 prognosis because they mount a stronger immunological response to their cancers. <sup>9</sup> In

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76 addition, surgeons may perform more extensive nodal dissections in healthier patients, 77 because they are judged more fit to tolerate a bigger operation or because their greater life expectancy prompts a more "aggressive" procedure, such as dissection down to the base 78 79 of the ileocolic vessel at its takeoff of the superior mesenteric artery for a right colectomy. 80 Associations between lymph node counts and survival may also be confounded by 81 provider factors. For example, higher lymph node examination rates may be correlated 82 with other provider attributes associated with improved survival after surgery for colon 83 cancer, including hospital volume and physician specialization (both surgeons and pathologists). <sup>10-16</sup> 84

Using data from the national Surveillance Epidemiology and End Results (SEER)-Medicare database, we performed a retrospective cohort study of patients undergoing resection for colon cancer. To better control for unmeasured confounding by patient factors, we examined relationships between lymph node counts and survival at the hospital level, rather than at the patient level. Specifically, we assessed whether hospitals with higher lymph node examination rates had better late survival rates than hospitals with lower rates, controlling for measurable patient and provider factors.

92

#### 93 Methods

<u>Subjects and databases.</u> For this study, we used the 1995-2005 national SEERMedicare linked database. As detailed elsewhere, these files provide a rich source of
information on Medicare patients included in SEER, a nationally representative collection of
population-based registries of all incident cancers from diverse geographic areas in the US. <sup>17</sup>
By the end of the study period, data from population-based cancer registries represented

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99 approximately 26% of the US population. For each Medicare patient in SEER, the SEER-

100 Medicare linked files contain 100% of Medicare claims from the inpatient, outpatient,

101 physician, home health, and hospice files.

102 From these files, we identified all patients aged 65 to 99 undergoing major resection for 103 colon cancer between 1995 and 2002. All Medicare patients with incident cases of these 104 cancers were identified by the appropriate cancer codes from the SEER files. Those patients 105 undergoing colectomy were identified from the Medicare Inpatient file using the appropriate 106 procedure codes from the International Classification of Diseases, version 9. Because lymph 107 node counts are less relevant for this population, we excluded patients with distant metastases 108 (stage IV disease). We also excluded the small proportion of patients who received 109 preoperative radiation therapy, which may confound lymph node counts.

110 Node examination rates. We identified all US hospitals at which SEER-Medicare

111 **patients underwent colectomy during the study period.** We then characterized each

112 hospital according to the proportion of patients in whom at least 12 lymph nodes were

113 examined (the standard endorsed by NQF), as determined from the appropriate field within the

114 Patient Entitlement and Diagnosis Summary File (PEDSF) from SEER. Hospitals were ranked

and sorted into four approximately evenly sized patient groups (quartiles). We repeated our

analysis after grouping hospitals by their median lymph node counts, rather than proportions

117 over 12. Because the two exposures measures were highly correlated (coefficient 0.78), results

118 from this sensitivity analysis were nearly identical to those of the baseline analysis and are not

119 presented here.

<u>Analysis</u>. Our primary outcome measure was mortality, determined at five years
from the date of resection or through December 31, 2005, which is the end of our follow-

122 up period. We used Cox proportional hazards models to examine relationships between 123 hospital node counts and mortality, adjusting for patient characteristics, censoring at the 124 end of the follow-up period. We used the patient as the unit of analysis, with the 125 exposure (node examination quartile) measured at the hospital level. We adjusted for age 126 group (65-69, 70-74, 75-79, 80-84, 85+), gender, race (black, non-black), year of 127 procedure, acuity of the index admission (elective, urgent, emergent), tumor location 128 (right, transverse, left, sigmoid colon), and patient comorbidities. We also adjusted for 129 tumor category (Tis, T1, T2, T3, T4). Comorbidities were identified using information 130 from the index admission and inpatient encounters from the preceding 6 months, based on methods described by Elixhauser et al.<sup>18</sup> Risk factors were assessed for colinearity, 131 132 over-fitting and interactions.

133 Although we subsequently stratified our results by tumor stage, we did not adjust 134 for this variable to avoid introducing bias into our baseline analysis. Hospitals that 135 examine more lymph nodes may appear to have more node-positive patients, even if their 136 patient populations are identical to those at hospitals examining fewer nodes. Risk 137 adjustment would artifactually reward those former hospitals for their "sicker" patients 138 and create a bias toward over-estimating the survival benefit of examining more nodes. As described elsewhere<sup>19</sup>, inpatient, outpatient and physician claims files were used to 139 140 identify patients receiving adjuvant chemotherapy, defined as therapy occurring within 6 141 months before or after surgery. We did not adjust for receipt of adjuvant chemotherapy 142 in our baseline analysis. Chemotherapy for node-positive patients is hypothesized to be 143 part of the causal pathway underlying potential relationships higher lymph node counts 144 and improved survival and thus not a true confounder. We did, however, adjust for

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145 provider characteristics potentially associated with improved late survival after cancer 146 surgery, including hospital teaching status, hospital volume, and surgeon volume. 147 Since patients admitted to the same hospital may have correlated outcomes, we used marginal survival models that accounted for clustering by hospital.<sup>20</sup> We first 148 149 assessed within-cluster correlations in patient failure times and derived robust variance-150 covariance estimators. These estimators were then incorporated into our multivariate 151 Cox proportional hazard models assessing relationships between hospital lymph node 152 examination rates and survival. All p-values are two-tailed. The institutional review 153 board of the University of Michigan approved the study protocol. 154 155 Results 156 Although age and gender did not vary markedly across hospital quartiles, 157 hospitals with highest proportions of patients with 12 or more lymph nodes tended to 158 treat fewer black patients and more patients admitted electively. As expected, 159 hospitals with the highest lymph node examination rates had a slightly lower proportion 160 of patients with stage II (node negative) disease (38% vs. 39% at hospitals with lowest 161 rates) and a higher proportion of patients with stage III disease (32% vs. 28%, 162 respectively). Although statistically significant, there were no clinically important 163 differences in the overall use of adjuvant chemotherapy (26% vs. 25% at hospitals 164 with the highest vs. lowest node rates, respectively). (Table 1) 165 Hospitals with the highest lymph node examination rates were more likely to be

167 likely to be high volume centers (43% vs. 20%, respectively). Surgeon volume did not
168 vary considerably across the hospital quartiles. (Table 1)

169 Before adjusting for potentially confounding variables, hospitals with the highest 170 node examination rates had higher survival rates after resection than hospitals with the 171 lowest rates. (Figure 1) Unadjusted 5-year survival probabilities for the two hospital 172 groups were 55% and 51%, respectively (p < 0.001). Relative to hospitals with low node 173 examination rates, the unadjusted hazard ratio of mortality associated with high lymph 174 node examination rates was 0.88 (95% CI, 0.84-0.93). After adjusting for confounding 175 patient and provider factors, however, hospital lymph node examination rates were no 176 longer associated with survival after surgery (adjusted hazard ratio, highest vs. lowest, 177 0.95, 95% CI 0.88-1.03). (Table 2) After adjusting for potential confounding by 178 tumor category, we found that depth of penetration of the tumor had no effect on 179 mortality rates (adjusted HR 0.94, 95% CI 0.87-1.02). In addition to the small 180 differences in survival between hospital groups at the extremes (hospital quartiles 1 181 and 4), we found no evidence of a "dose-response" effect in the intermediate quartile 182 comparisons. Though none of these differences were statistically significant, 183 hospital quartile 2 had slightly worse survival than quartile 1 (lowest node rates); 184 hospital quartile 4 (highest node rates) had slightly worse survival than quartile 3. 185 (Table 2)

Table 2 also summarizes relationships between survival and hospital lymph node
examination rates in various patient subgroups. As suggested in previous studies,
adjusted hazard ratios of mortality between the hospitals with highest nodes counts and
lowest node counts were lowest among patients with stage II disease (adjusted HR 0.85,

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95% CI 0.74-0.96), compared with those for stage I disease (adjusted HR 0.92, 95% CI
0.78-1.09) and stage III disease (adjusted HR 0.98, 95% CI 0.86-1.11). Patient age and
tumor location were not important modifiers of the relationship between hospital lymph
node examination rates and late survival after surgery. (Table 2)

194

### 195 **Discussion**

196 Our study raises questions about the importance of examining a large number of 197 lymph nodes in patients with colon cancer. Using SEER-Medicare data, we profiled 198 hospitals according to how frequently they achieved the 12-node minimum suggested by 199 many experts and then assessed late survival according to this measure. In addition to 200 reducing risks of patient selection bias within hospitals, comparison at the hospital level 201 most directly simulates survival differences that would be observed if lymph node counts 202 were used as a hospital quality indicator for colon resections. Although hospitals with 203 high lymph node counts had higher observed survival rates than hospitals with low 204 counts, absolute survival differences were considerably smaller than previously estimated 205 by patient-level analyses. Moreover, these survival differences largely disappeared after 206 adjusting for confounding patient and provider characteristics.

207 Our main findings are consistent with one previous hospital-level analysis, 208 which similarly failed to detect a survival benefit associated with higher node counts 209 in a small cohort of patients undergoing colectomy in Canada.<sup>5</sup> However, our 210 findings differ from those of most previous studies, all based on patient-level 211 comparisons, which have suggested that more extensive nodal examination improves 212 patient outcomes.<sup>1-3, 21</sup> A recent systematic analysis summarized data from 17 studies

examining the impact of lymph node counts on survival after colectomy for colon 213 cancer.<sup>4</sup> It included population-based studies, single institution retrospective cohort 214 215 studies, and nested cohort studies which included prospectively gathered data from 216 clinical trials. These studies were too heterogeneous with regards to study populations, 217 outcome measures and node cut-points to allow for formal meta-analysis. However, 218 survival differences between patients with high and low lymph node counts were 219 apparent in most studies. In several, patients in whom the greatest numbers of nodes 220 were examined had hazard ratios of mortality below 0.70, relative to patients with the 221 fewest nodes.

222 Although their study populations and analytic methods vary widely, previous 223 studies share several limitations that may lead to overestimating the true benefits 224 associated with more extensive lymph node dissection. First, as described earlier, prior 225 studies assessing relationships between lymph node counts and survival may have failed 226 to fully account for confounding patient characteristics. In this study, higher node counts 227 were associated with fewer comorbidities, lower admission acuity, and right-sided colon 228 cancers. Although differences in case mix were relatively small in magnitude, they were 229 sufficient to explain small differences in survival rates associated with hospital lymph 230 node examination rates. Moreover, differences in measurable risk factors increase the 231 likelihood of additional confounding by unmeasured variables. It is plausible that 232 surgeons perform more extensive nodal dissections in healthier patients, because they are 233 judged more fit to tolerate potential complications or because their greater life expectancy 234 prompts the surgeons to be more "aggressive." By focusing on hospital-level 235 comparisons, our study was designed to minimize patient selection bias within hospitals.

Of course, this design does not deal with confounding by unmeasured differences in case mix across hospitals. To this extent, our study no doubt underestimates the extent to which relationships between hospital lymph node examination rates and survival are confounded by patient characteristics.

240 Second, previous studies have failed to fully account for confounding provider 241 attributes. In our study, hospitals with higher lymph node counts were more likely to be 242 teaching hospitals and had higher hospital procedure volumes. A growing number of 243 studies suggest that higher hospital volumes are associated with improved survival after cancer surgery. <sup>10, 16</sup> Some would argue that obtaining more lymph nodes may be part of 244 the causal pathway underlying apparent relationships between volume and late survival, 245 246 and thus that we should not have treated volume as a confounder in this study (and 247 adjusted for it). However, our previous analyses suggest that volume and late survival 248 associations are largely independent of lymph nodes counts, implying that volume exerts its effects by a myriad of other processes.<sup>10</sup> 249

250 Finally, to avoid biases associated with cancer upstaging at hospitals with high 251 node examination rates, our study focused primarily on overall survival among all 252 patients with non-metastatic colon cancer, not stage-specific survival. However, 253 consistent with findings from previous studies, we found that node examination rates 254 appeared to be most important in patients with stage II (node-negative) disease. This 255 result is not surprising. Hospitals that examine more lymph nodes will tend to find more patients with node-positive disease <sup>22</sup> and more accurately classify node-negative 256 257 patients. In contrast, more patients classified as stage II at hospitals with low node 258 examination rates may truly be node-positive and experience lower survival. To the

extent that it improves cancer staging, higher node examination rates may be valuable in
providing stage I or II patients with more accurate prognoses. However, a true survival
benefit associated with higher lymph node counts should be evident in improved survival
among all patients undergoing colon cancer resection. We failed to detect such benefits
in our analysis.

264 Our study has several limitations. First, our study was limited to Medicare 265 patients over 65 years of age, who represent approximately two-thirds of patients with new diagnoses of colon cancer.<sup>23</sup> The use of SEER-Medicare data, compared to SEER 266 267 data alone, allowed us to account for patient comorbidities, admission acuity, and 268 provider attributes, all important confounders of relationships between lymph node 269 counts and survival after cancer surgery. In stratified analyses, we found no significant 270 evidence that patient age is an important modifier of relationships between survival and 271 hospital node examination rates. Nonetheless, the generalizability of our findings to patients under 65 years of age is unknown. 272

273 Second, our hospital-level comparisons essentially ignore wide variations in the 274 number of nodes examined among patients within hospitals. It is unlikely that such 275 variation is primarily attributable to differences of quality of care within hospitals. At 276 most hospitals, surgical specimens from gastrointestinal resections are evaluated by a 277 small number of pathologists, if not by a single physician. Although there may be 278 variation in surgical technique among surgeons at the same hospital, our conclusions 279 were unaltered when we assessed lymph node examination rates at the surgeon level 280 rather than at the hospital level. For these reasons, we believe that within-hospital

variation in the numbers of lymph nodes examined primarily reflects patient factorsrelated to tumor biology and immunology, not quality of care.

283 Although our study documents wide variation in node examination rates 284 across hospitals, it does not provide insights into why some hospitals are more 285 successful than others in this regard. Although we cannot rule out variation in 286 surgical technique, practices related to processing and evaluation of surgical 287 specimens in the pathology department may also be responsible. Recent studies 288 have suggested that wide variation in the quality of gross specimen preparation and node extraction, tasks which are often performed by technicians.<sup>24</sup> Pathologists 289 may also vary in their skill or diligence in identifying nodes. <sup>25</sup> A better 290 291 understanding of mechanisms underlying variation in hospital node examination rates with colon cancer would be essential for increasing node counts. 292

293 Using lymph node counts as a hospital quality indicator is gaining momentum 294 from stakeholders in the health care community. For instance, as part of their pay for 295 performance programs, several large private payers have already begun to hold providers accountable for recovering at least 12 nodes following resection for colon cancer.<sup>7</sup> Our 296 297 study also suggests that the potential gains in patient outcomes associated with 298 improvements in this process of care may be smaller than many believe. Further studies 299 based on datasets with more clinical detail would be useful for confirming or refuting our 300 findings, and for identifying more effective levers for improving quality of care in 301 patients with colon cancer.

- 302 Disclosures
- 303 Author contributions. Dr. Birkmeyer had full access to all of the data in the study and takes
- 304 responsibility for the integrity of the data and the accuracy of the data analysis.
- 305 Study concept and design: Birkmeyer, Hollenbeck, and Wong
- 306 Acquisition of data: Birkmeyer
- 307 Analysis and interpretation of data: Birkmeyer, Ji, Baser, and Wong
- 308 Drafting of the manuscript: Birkmeyer, Wong
- 309 Critical revision of the manuscript for important intellectual content: Birkmeyer, Baser,
- 310 Hollenbeck, Morris, and Wong
- 311 Statistical analysis: Ji, Baser
- 312 Obtained funding: Birkmeyer
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- 314 Study supervision: Birkmeyer, Baser
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# 389 Figure legend

- **Figure 1.** Kaplan-Meier plot describing 5-year survival among patients undergoing
- 391 resection for colon cancer, according to hospital lymph node examination rates.
- Based on data from the SEER-Medicare linked database, 1995-2005.

393 **Table 1.** Characteristics of patients undergoing resection for colon cancer, according to

- hospital lymph node examination rates. Based on data from the 1995-2005 SEER-Medicare
- 395 database.

|  | Hospital lymph node examination |                             |                             |                             |             |  |
|--|---------------------------------|-----------------------------|-----------------------------|-----------------------------|-------------|--|
|  | rates (Quartile)                |                             |                             |                             |             |  |
|  | 1 <sup>st</sup><br>quartile     | 2 <sup>nd</sup><br>quartile | 3 <sup>rd</sup><br>quartile | 4 <sup>th</sup><br>quartile | p-<br>value | Hazard ratio of<br>mortality associated<br>with variable (95%<br>CI) |
| Number of patients   | 7526                            | 7633                        | 7513                        | 7953                        |             |  |
| Number of hospitals  | 233                             | 225                         | 153                         | 289                         |             |  |
| Median number of<br>lymph nodes examined                           | 6                               | 8                           | 10                          | 13                          |             |  |
| Mean of lymph nodes examined                                       | 5.9                             | 7.7                         | 9.5                         | 14.9                        |             |  |
| Proportion of patients<br>with at least 12 lymph<br>nodes examined | 16.2%                           | 31.2%                       | 43.4%                       | 60.7%                       |             |  |
|  |                                 |                             |                             |                             |             |  |
| Patient characteristics  |                                 |                             |                             |                             |             |  |
| Age (%)  |                                 |                             |                             |                             | 0.46        |  |
| 65-69  | 14.6                            | 15.0                        | 15.2                        | 15.8                        |             | 1  |
| 70-74  | 21.1                            | 21.5                        | 21.1                        | 21.4                        |             | 1.28(1.19,1.37)  |
| 75-79  | 23.9                            | 24.4                        | 24.9                        | 24.8                        |             | 1.59(1.49,1.69)  |
| 80-84  | 21.3                            | 20.3                        | 20.4                        | 20.1                        |             | 2.20(2.06,2.34)  |
| 85+  | 19.1                            | 18.8                        | 18.4                        | 18.0                        |             | 3.44(3.23,3.66)  |
| Gender (% female)  | 56.4                            | 56.4                        | 57.1                        | 56.1                        | 0.71        | 0.93(0.90,0.96)  |
| Race (% black)   | 8.1                             | 7.0                         | 9.4                         | 5.4                         | <.001       | 1.17(1.10,1.24)  |
| Admission acuity (%)   |                                 |                             |                             |                             | <.001       |  |
| Elective   | 53.8                            | 57.9                        | 60.2                        | 63.7                        |             | 1  |
| Urgent   | 25.2                            | 22.8                        | 17.1                        | 17.9                        |             | 1.62(1.56,1.69)  |
| Emergent   | 21.0                            | 19.3                        | 22.7                        | 18.4                        |             | 2.28(2.19,2.37)  |
| Comorbidity (% 2+)   | 64.2                            | 63.1                        | 64.2                        | 61.7                        | 0.003       | 1.99(1.91,2.06)  |
| Tumor location   |                                 |                             |                             |                             | <.001       |  |
| Right  | 54.9                            | 55.9                        | 56.4                        | 58.7                        |             | 1  |
| Transverse   | 9.9                             | 9.9                         | 9.9                         | 9.7                         |             | 1.04(0.98,1.11)  |

| Left                            | 8.8  | 9.8  | 9.7  | 9.2  |       | 1.06(0.99,1.13)  |
|---------------------------------|------|------|------|------|-------|------------------|
| Sigmoid                         | 26.4 | 24.4 | 24.1 | 22.5 |       | 0.92(0.88,0.96)  |
| Tumor stage (modified AJCC)     |      |      |      |      | <.001 |                  |
| Stage 0/I                       | 32.5 | 31.5 | 30.7 | 29.9 |       | 1                |
| Stage II                        | 39.4 | 40.2 | 39.5 | 38.3 |       | 1.41(1.35,1.47)  |
| Stage III                       | 28.1 | 28.3 | 29.8 | 31.7 |       | 2.26(2.16,2.36)  |
| Adjuvant therapy                |      |      |      |      |       |                  |
| Chemotherapy                    | 25.2 | 27.8 | 25.2 | 26.4 | <.001 | 0.94(0.90, 0.97) |
|                                 |      |      |      |      |       |                  |
| <b>Provider characteristics</b> |      |      |      |      |       |                  |
| Teaching (%)                    | 32.5 | 38.2 | 49.4 | 58.3 | <.001 | 0.98(0.95,1.02)  |
| Hospital procedure volume (%)   |      |      |      |      | <.001 |                  |
| Low                             | 45.7 | 38.0 | 22.2 | 27.7 |       | 1                |
| Medium                          | 34.8 | 31.4 | 38.0 | 29.2 |       | 0.93(0.89,0.97)  |
| High                            | 19.5 | 30.5 | 39.8 | 43.0 |       | 0.90(0.86,0.93)  |
| Surgeon procedure volume (%)    |      |      |      |      | <.001 |                  |
| Low                             | 31.7 | 35.4 | 34.6 | 31.8 |       | 1                |
| Medium                          | 32.7 | 33.3 | 32.6 | 34.6 |       | 1.02(0.97,1.07)  |
| High                            | 35.6 | 31.3 | 32.8 | 33.6 |       | 1.08(1.02,1.13)  |

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|                | Adjusted hazard ratio of mortality according to hospital<br>node examination rates, (95% CI) |  |  |  |  |  |  |
|----------------|--|--|--|--|--|--|--|
|                | Referent group:<br>1 <sup>st</sup> quartile<br>(lowest lymph<br>node counts)                 | 2 <sup>nd</sup> quartile<br>versus 1 <sup>st</sup><br>quartile | 3 <sup>rd</sup> quartile<br>versus 1 <sup>st</sup><br>quartile | 4 <sup>th</sup> quartile<br>versus 1 <sup>st</sup><br>quartile |  |  |  |
| All patients   | 1.0  | 1.04 (0.96,1.12)   | 0.94 (0.87,1.02)   | 0.95 (0.88,1.03)   |  |  |  |
|                |  |  |  |  |  |  |  |
| Age            |  |  |  |  |  |  |  |
| 65-69          | 1.0  | 1.02 (0.82,1.26)   | 0.92 (0.73,1.15)   | 0.90 (0.72,1.11)   |  |  |  |
| 70-74          | 1.0  | 1.00 (0.84,1.21)   | 0.93 (0.77,1.13)   | 0.96 (0.79,1.16)   |  |  |  |
| 75+            | 1.0  | 1.05 (0.96,1.14)   | 0.95 (0.87,1.04)   | 0.95 (0.87,1.04)   |  |  |  |
| Stage          |  |  |  |  |  |  |  |
| 0/I            | 1.0  | 1.04 (0.88,1.21)   | 0.88 (0.75,1.04)   | 0.92 (0.78,1.09)   |  |  |  |
| II             | 1.0  | 0.99 (0.88,1.13)   | 0.94 (0.82,1.06)   | 0.85 (0.74,0.96)   |  |  |  |
| III            | 1.0  | 1.06 (0.93,1.20)   | 0.91 (0.80,1.04)   | 0.98 (0.86,1.11)   |  |  |  |
| Tumor location |  |  |  |  |  |  |  |
| Right          | 1.0  | 1.00 (0.90,1.12)   | 0.92 (0.82,1.03)   | 0.94 (0.84,1.06)   |  |  |  |
| Left/sigmoid   | 1.0  | 1.12 (0.98,1.29)   | 0.97 (0.84,1.12)   | 0.99 (0.85,1.14)   |  |  |  |

**Table 2.** Association between hospital node examination rates and late survival after

398 colectomy for colon cancer, with adjustment for patient and provider characteristics.

399 Adjusted for the following patient and provider characteristics: age, sex, race, year of

400 procedure, admission acuity, patient comorbidities, location of tumor, hospital teaching

401 status, hospital volume, and surgeon volume

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